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# Frequency-Based Parameterization of Semi-Empirical Models for State-of-Health Estimation in Lithium-Ion Batteries

Colin J. Chu<sup>a</sup>, Sai Thatipamula<sup>b</sup>, and Simona Onori<sup>\*b</sup>

<sup>a</sup>The Nueva School, San Mateo, CA 94403, USA

<sup>b</sup>Department of Energy Science & Engineering, Stanford University, Stanford, CA 94305, USA

\*Corresponding author: [sonori@stanford.edu](mailto:sonori@stanford.edu)

## Abstract

The accurate estimation of state-of-health (SoH) in lithium-ion batteries is crucial for enabling safe and improved battery usage. Yet the complexity of electrochemical aging processes within the cell and their dependence on differing operating conditions make this task challenging. This paper proposes a novel SoH estimation method using equivalent circuit model (ECM) parameters identified based on electrochemical impedance spectroscopy and regression analysis. Six different ECM architectures are calibrated and compared based on quantitative metrics including root mean square error, Akaike information criterion, and Bayesian information criterion. Based on this analysis, the second-order constant phase element model is chosen and its parameters are used to perform three supervised tree-based regressions and one supervised linear regression. The random forest regression model is found to be the most accurate in estimating capacity-based SoH and a rigorous correlation analysis is conducted to examine the relationship between model parameters and SoH. This study utilizes data collected from an aging campaign of 22 5Ah nickel-manganese-cobalt lithium-ion cells and demonstrates the effectiveness of ECMs and regression models for SoH estimation.

**Keywords:** Lithium-ion battery, state-of-health, modeling and estimation, electrochemical impedance spectroscopy, equivalent circuit models.

## 1. Introduction

Lithium-ion (Li-ion) batteries deliver long-term efficiency and offer unique characteristics, including high gravimetric and volumetric energy density [1–3]. Li-ion batteries exhibit additional properties, including low self-discharge rates, extended cycle lives, and fast

charging capabilities [4–7]. Nonetheless, Li-ion cells degrade due to several mechanisms, including solid electrolyte interface growth, lithium plating, and particle cracking [8, 9]. In addition, further degradation can be triggered by transition metal dissolution and the decomposition of the electrolyte [10, 11]. Accurate modeling of capacity-based state-of-health (SoH) is necessary for the optimization of battery management system (BMS) performance [12–14]. Capacity-based SoH is defined by:

$$SoH(t) = \frac{Q(t)}{Q_0} \cdot 100\% \quad (1)$$

where  $SoH(t)$  is expressed as a percentage of the measured capacity at time  $t$ ,  $Q_t$ , and the nominal capacity when the battery is new,  $Q_0$  [15]. Existing SoH estimation methods are constrained by the complexity of the electrochemical degradation mechanisms they can model, as well as limitations in robustness and accuracy when implemented in BMS applications [16, 17].

In this work, we use electrochemical impedance spectroscopy (EIS) data, which measures the impedance, or frequency-dependent opposition of the battery to current flow, to estimate SoH [18]. The EIS data used in this paper, as described in the Data section, provides key insights into the electrochemical processes in Li-ion batteries such as charge transfer, ion diffusion, and double-layer capacitance [19–21].

While previous literature has focused on equivalent circuit model (ECM)-based methods for real-time SoH estimation, this work proposes a novel approach that leverages regression techniques to effectively develop a semi-empirical pipeline for SoH estimation [22, 23]. We propose a reliable and comprehensive method to predict capacity-based SoH by extracting the parameters of the best-fitting ECM, which was determined via a comparative analysis. ECMs simulate the electrochemical response of the cell by modeling it using electrical components [25–28]. However, under certain dynamic conditions, ECMs may exhibit limitations due to their inability to account for the non-linear characteristics and non-idealities of the cells [29, 30].

Previous work has also included the use of single-particle model parameters, which can be uniquely determined and fit using frequency-domain data for reliable time-domain simulations [31]. Subsequent research has refined the parameterization of physics-based models across states of charge using grouped-variable approaches, streamlining interpretation while preserving electrochemical relevance. These refinements also incorporate double-layer capacitance into the models to replicate the semicircle observed in Nyquist plots [32]. Recently, hybrid manual-automatic feature extraction techniques and light-weight convolutional neural network (CNN) architectures have also been proposed to deliver fast, computationally efficient capacity estimation methods that remain robust across varying cell chemistries and measurement uncertainties [33, 34]. In addition, recent studies have advanced data-driven SoH estimation, including LSTM-based frameworks for degradation pattern recognition, physics-informed neural networks embedding incremental

capacity peak constraints, and hybrid CNN–TCN–LSTM fusion architectures optimized via Bayesian methods for multi-feature learning [35–37]. Finally, Giazitzis et al. introduce a TinyML framework for SoH estimation from EIS data, though it is limited to only five cells under fixed conditions without temperature variation and does not benchmark different ECM families [38]. In contrast, our work uses 22 cells aged under varied conditions, systematically compares six ECMs, and applies regression with correlation analysis to deliver model-specific interpretability of feature contributions to SoH.

In this work, three supervised regression models are individually trained and tested, using the identified parameters from the most accurate and efficient ECM as input features. Regression models consider complex processes and tolerate noisy data; in addition, many models exhibit high interpretability in providing key insights, such as feature importance [39–42]. Nonetheless, it is important to acknowledge that regression models alone may also be oversensitive to specific data, resulting in overfitting [43,44]. Thus, the combination of ECMs and regression models provides an efficient and interpretable approach to SoH estimation.

The paper structure is as follows: the Experimental Data section discusses the EIS data used in this work; the Methods section details the parametrization and fitting of the different ECMs, the selection of the most parsimonious ECM (i.e., the model that best balances accuracy with simplicity), and the construction of the supervised regression models to estimate SoH; the Results section reports the best model architecture and fitted parameters; and the Conclusions section summarizes the novelty of the research and proposes potential applications and future directions of work.

## 2. Experimental Data

The experimental data used in this work is discussed rigorously in Khan et al [45]. EIS data was collected from an aging campaign performed on 5Ah NMC (nickel-manganese-cobalt 1-1-1)/graphite prismatic Li-ion batteries using a frequency response analyzer and potentiometer/galvanostat by KRI, Inc. The dataset consists of 22 cells labeled S1 through S24, with S2 and S18 excluded due to evaluation errors. Cells S1, S3, S4, S8, and S9 were calendar-aged and were neither charged nor discharged during the experimental campaign. The remaining 17 cells were cycle-aged. Calendar-aged versus cycle-aged cells did not exhibit statistically significantly different results.

## 3. Methods

In this work, we compare ECMs based on their ability to inform Li-ion cells' capacity-based SoH. As shown in Figure 1, the methods used to achieve this goal involve several steps. After the EIS data is collected and processed, various ECMs are fit to the data to extract electrochemical parameters. The most appropriate ECM is then selected based on fitting

accuracy and efficiency. Finally, data-driven SoH estimation models are built and compared using the extracted parameters, including feature attribution and error analysis.

The above pipeline leads to a framework to estimate SoH of Li-ion cells using frequency-domain data. This work makes three key contributions. First, this study conducts a direct and systematic evaluation of six ECMs across three distinct circuit families (Randles, ZARC, and Transmission Line) [22, 47]. Second, while prior work has leveraged deep learning architectures, such as CNN-BiLSTM (Bidirectional Long Short-Term Memory) or GPR (Gaussian Process Regression)-Elman networks, for SoH prediction using modified Randles or ZARC parameters, this study focuses on interpretable supervised regressors (decision tree, random forest, gradient boosting, and linear regression) to allow for feature attribution [22, 48]. Third, this work incorporates rigorous interpretability by applying SHapley Additive exPlanations (SHAP) analysis to quantify the contribution of each ECM-derived parameter to the SoH prediction, enabling feature-level insights often absent in prior studies. Together, these contributions establish a scalable and interpretable framework for SoH estimation.

### 3.1 Model Parameterization and Fitting

Six ECMs were fitted with the EIS data: two types of second-order Randles circuits (e.g. 1W Randles, 2W Randles), three ZARC circuits of different orders (e.g. 1 ZARC, 2 ZARC, 3 ZARC), and one second-order Transmission Line model. All three families of models have been explored independently in Li-ion batteries in past literature [50–52], but a direct comparative analysis between the three ECMs is missing. The circuit diagrams of the ECMs used in this study are shown in Figure 2.

Model	Equation
1W Randles	$Z(\omega)_{1W \text{ Randles}} = j\omega L_0 + R_0 + \left(\frac{1}{R_1} + j\omega C_1\right)^{-1} + \left(\frac{1}{R_2} + j\omega C_2\right)^{-1} + \frac{W_0}{\sqrt{\omega}}(1-j)$
2W Randles	$Z(\omega)_{2W \text{ Randles}} = j\omega L_0 + R_0 + \left(\frac{1}{R_1} + j\omega C_1\right)^{-1} + \left(\frac{1}{R_2 + W_1} + j\omega C_2\right)^{-1} + \frac{W_0}{\sqrt{\omega}}(1-j)$
1 ZARC	$Z(\omega)_{1 \text{ ZARC}} = j\omega L_0 + R_0 + \frac{R_1}{1+(j\omega\tau_1)^{\alpha_1}} + \frac{W_0}{\sqrt{\omega}}(1-j)$
2 ZARC	$Z(\omega)_{2 \text{ ZARC}} = j\omega L_0 + R_0 + \frac{R_1}{1+(j\omega\tau_1)^{\alpha_1}} + \frac{R_2}{1+(j\omega\tau_2)^{\alpha_2}} + \frac{W_0}{\sqrt{\omega}}(1-j)$
3 ZARC	$Z(\omega)_{3 \text{ ZARC}} = j\omega L_0 + R_0 + \frac{R_1}{1+(j\omega\tau_1)^{\alpha_1}} + \frac{R_2}{1+(j\omega\tau_2)^{\alpha_2}} + \frac{R_3}{1+(j\omega\tau_3)^{\alpha_3}} + \frac{W_0}{\sqrt{\omega}}(1-j)$
2 TL	$Z(\omega)_{2 \text{ TL}} = j\omega L_0 + R_0 + \sqrt{\frac{R_{ion1}}{Q_{s1}(i\omega)^{\alpha_1}}} \coth(\sqrt{Q_{s1}(i\omega)^{\alpha_1}} R_{ion1}) + \sqrt{\frac{R_{ion2}}{Q_{s2}(i\omega)^{\alpha_2}}} \coth(\sqrt{Q_{s2}(i\omega)^{\alpha_2}} R_{ion2}) + \frac{W_0}{\sqrt{\omega}}(1-j)$

Table 1: The equations of the six ECMs used in this study. The bolded variables denote ECM parameters that were optimized during fitting.

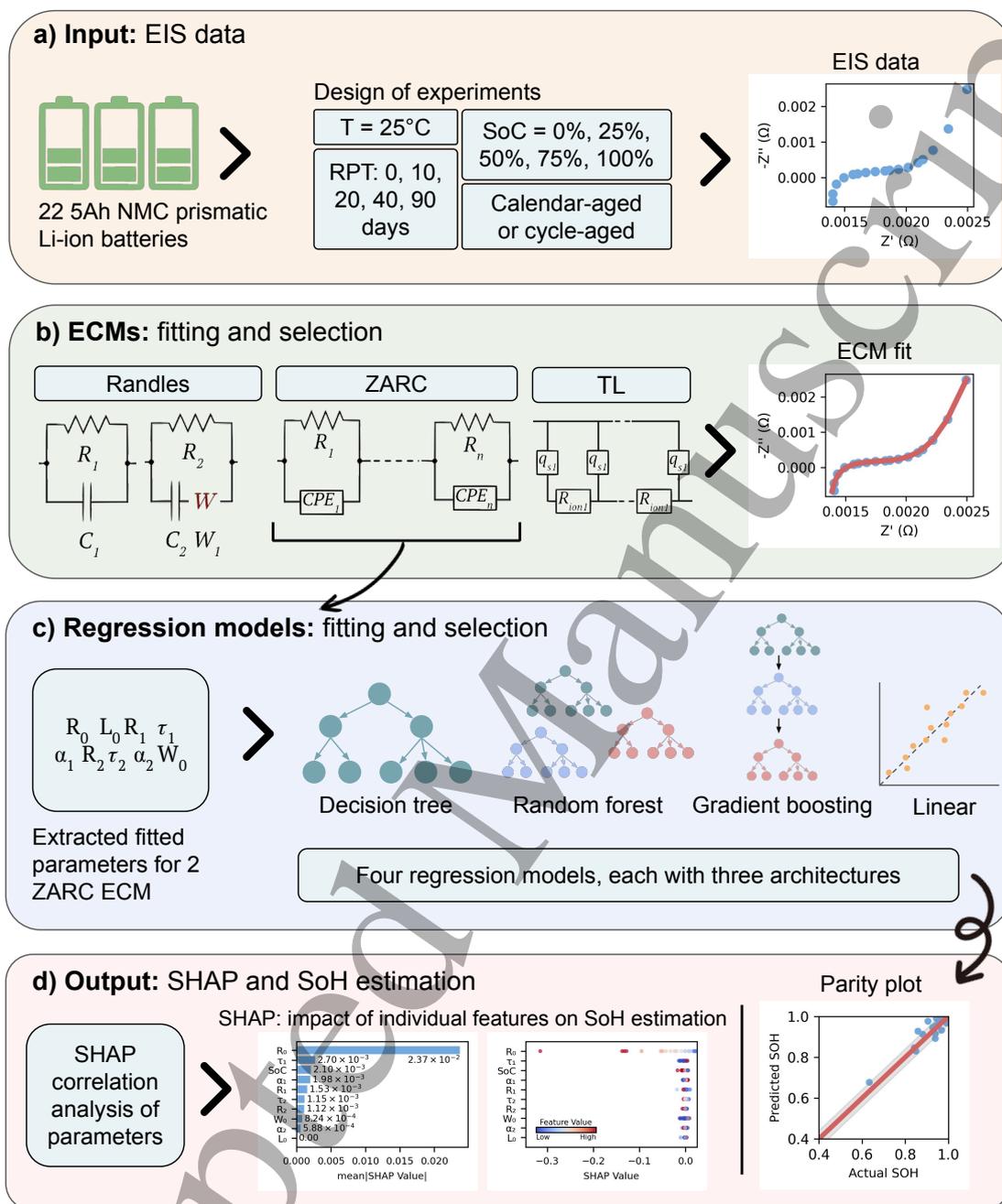


Figure 1: SoH estimation framework. (a) EIS data is collected from 22 NMC Li-ion cells under varied SoC and aging conditions, (b) multiple ECMs are fitted to extract electrochemical parameters, (c) regression models are trained on selected parameters to predict battery SoH, and (d) SHAP analysis quantifies feature importance, and model performance is evaluated.

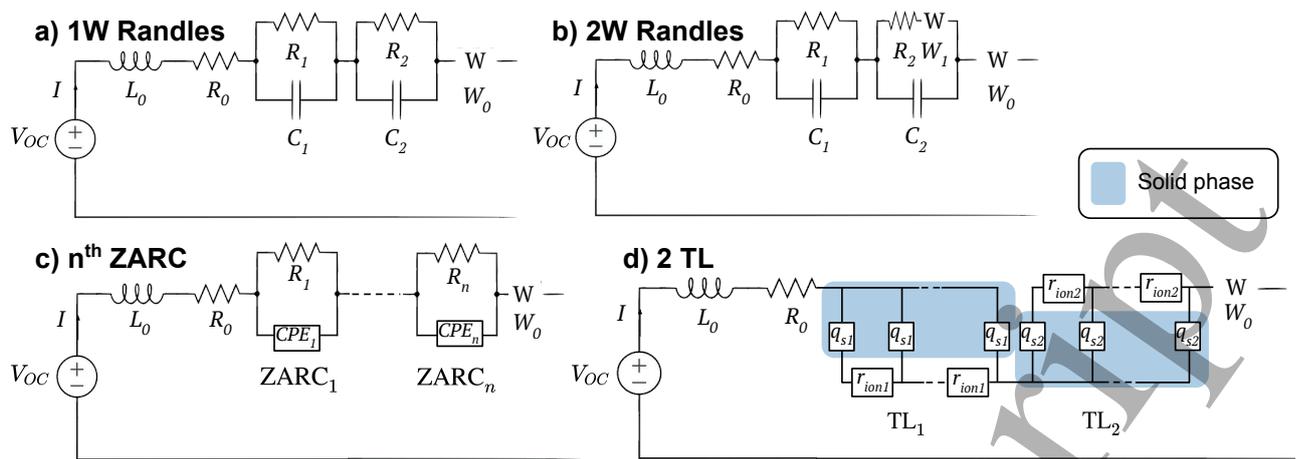


Figure 2: In this study, six different ECMs are used: 1W Randles, 2W Randles, 1 ZARC, 2 ZARC, 3 ZARC, and 2 TL. This figure includes the circuit diagrams of the a) 1W Randles, b) 2W Randles, c)  $n^{\text{th}}$  ZARC, and d) 2 TL ECMs.

Using ECM parameters fitted from EIS data is more advantageous than estimating SoH directly from raw EIS data. ECM fitting reduces the dimensionality of the problem by compressing the full impedance data into a set of physically interpretable features. This reduces the risk of overfitting. In addition, given the interpretability of ECM parameters, the resulting model is able to more clearly explain the underlying degradation mechanisms.

The Randles circuit is a widely recognized model that fits EIS data by simulating behaviors at the electrode interface [53, 54]. This paper specifically evaluates a second-order Randles circuit with one Warburg component, referred to as the 1W Randles circuit, and a second-order Randles circuit with two Warburg components, referred to as the 2W Randles circuit. The 1W Randles circuit consists of the ohmic resistor, inductor, two resistor-capacitor (RC) components, and a Warburg element. Each part can be explained in terms of what electrochemical phenomenon it models and how it contributes to the EIS fit. The inductance element ( $L_0$ ) models opposition to changes in current flow and appears as a vertical line in the positive imaginary region. The resistor ( $R_0$ ) models resistance of the electrolyte, electrodes, and current collectors and appears as the EIS data's intercept with the real axis. The first resistor-capacitor element ( $R_1C_1$ ) models the solid-electrolyte interface (SEI) while the second resistor-capacitor element ( $R_2C_2$ ) models charge-transfer resistance and double-layer capacitance [55, 56]. Both RC elements appear as semi-circles. Finally, the Warburg element ( $W_0$ ) models the diffusion process and mass-transfer limitations, appearing as a 45-degree tail at low frequencies. The 2W Randles circuit consists of an additional Warburg element in series with the resistor of the second RC element [57].

In contrast to the 1W Randles circuit, the 2W Randles configuration not only captures a semi-infinite diffusion process but also a limited or localized diffusion process as well [58]. The equations for the impedance of the two Randles circuits are shown in Table 1, where  $j$  is the imaginary unit  $\sqrt{-1}$ ;  $\omega$  is the angular frequency [rad/s];  $L_0$  is the inductance

1  
2 [H];  $R_0$ ,  $R_1$ , and  $R_2$  are resistors [ $\Omega$ ];  $C_1$  and  $C_2$  are capacitors [ $F$ ]; and  $W_0$  and  $W_1$  are 152  
3 Warburg coefficients [ $\Omega/s^{1/2}$ ] [59, 60]. 153

4  
5 The second and more nuanced ECM examined is the ZARC circuit, which is widely used 154  
6 to represent the non-ideal behavior of the double-layer capacitance and charge transfer 155  
7 in Li-ion batteries [61]. The ZARC model is a physically informed extension of the 156  
8 classic parallel resistor-capacitor (RC) circuit, designed to capture real-world imperfections 157  
9 in interfacial processes at the electrode–electrolyte boundary. This circuit introduces 158  
10 a constant phase element (CPE), which is defined by two parameters. First,  $\tau$  is the 159  
11 characteristic time constant, physically interpreted as the product of the charge-transfer 160  
12 resistance and a generalized capacitance. It determines the characteristic frequency range 161  
13 over which the interfacial process occurs. Second,  $\alpha$  is the CPE exponent, bounded 162  
14 between 0 and 1. It quantifies the degree of deviation from ideal capacitive behavior: 163  
15  $\alpha = 1$  corresponds to an ideal capacitor with uniform surface and perfect time response 164  
16 while  $\alpha = 0$  behaves like a pure resistor. 165

17  
18 A ZARC element is a resistor in parallel with a CPE, and has the effect of depressing 166  
19 the semicircular shape that an RC element would produce in a Nyquist plot, as shown 167  
20 in Figure 3. In this paper, the first-order, second-order, and third-order ZARC circuits 168  
21 (referred to as 1 ZARC, 2 ZARC, and 3 ZARC, respectively) are evaluated. The equations 169  
22 for all three ZARC circuits are shown in Table 1. 170

23  
24 The third ECM family is the Transmission Line (TL), which models the electrochemical 171  
25 behavior of electrolyte-filled porous electrodes [62–64]. In this work, we focus on the 172  
26 second-order TL model (2 TL), which consists of a repeating unit of two physically derived 173  
27 elements:  $q_s$  and  $R_{ion}$  [65, 66]. First,  $q_s$  models the charge transfer processes at the interface 174  
28 between the electrode's solid matrix and the liquid electrolyte. It typically behaves like a 175  
29 constant phase element (CPE), reflecting the distributed nature of electrochemical reactions 176  
30 across the porous surface. The CPE is characterized by  $Q_s$ , the charge defined as the sum 177  
31 of all  $q_s$  elements, and  $\alpha$ , the exponent of the CPE, or the degree of non-ideality in the 178  
32 interfacial charge transfer.  $R_{ion}$ , which is defined as the sum of all  $r_{ion}$  elements, represents 179  
33 the resistance to movement within the liquid electrolyte occupying the pore volume. This is 180  
34 a purely resistive component and is directly related to the ionic conductivity and tortuosity 181  
35 of the electrolyte pathways. Higher  $R_{ion}$  indicates either longer or more convoluted ionic 182  
36 paths, or lower electrolyte conductivity [67, 68]. 183

37  
38 In the 2 TL structure, multiple  $R_{ion}$  and  $q_s$  units may be connected in series, mimicking 184  
39 a finite section of a porous electrode where each pair represents a small slice of the pore 185  
40 depth. This ladder-like structure captures the distributed nature of ion transport and 186  
41 interfacial reaction across the electrode depth. The resulting impedance equation for the 187  
42 2 TL model, provided in Table 1, is derived from solving the equivalent circuit based 188  
43 on Ohm's law and the current-voltage relationships of CPEs, reflecting a physics-based 189  
44 approximation of ion diffusion and charge transfer in porous media [65]. 190

45  
46 The Python-based package `impedance.py` was utilized for the fitting of all six ECMs [69]. 191

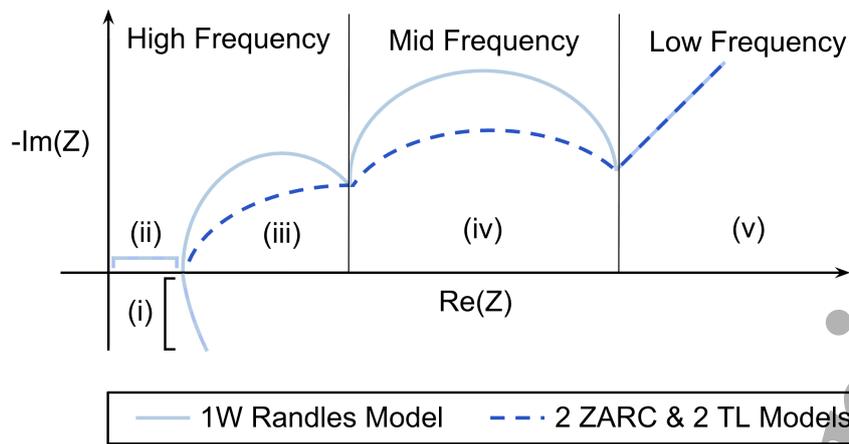


Figure 3: Nyquist plot demonstrating the theoretical impact of individual ECM components: (i) inductance element  $L_0$ ; (ii) ohmic resistance element  $R_0$ ; (iii)  $R_1-C_1$  component for 1W Randles Model,  $ZARC_1$  component for 2 ZARC and 2 TL Models; (iv)  $R_2-C_2$  component for 1W Randles Model,  $ZARC_2$  component for 2 ZARC and 2 TL Models; and (v) Warburg element  $W_0$ .

`impedance.py` is capable of fitting EIS data to custom circuits composed of up to 14 unique pre-defined elements. The package utilizes `curve_fit` from `scipy.optimize` function, which is a non-linear least squares regression fitting method. As no bounds were defined in these methods, the `curve_fit` function defaulted to using the Levenberg-Marquardt optimization algorithm to minimize the objective function, defined as:

$$\chi^2 = \sum_{n=1}^N [Z'_{EIS}(\omega_n) - Z'_{ECM}(\omega_n)]^2 + [Z''_{EIS}(\omega_n) - Z''_{ECM}(\omega_n)]^2, \quad (2)$$

where  $Z'_i$  is the real component of impedance and  $Z''_i$  is the imaginary component of impedance, where the  $i$  subscript distinguishes between the experimental EIS data and the ECM fit, respectively [70].  $N$  is the number of frequencies at which the EIS measurements were taken; in the case of the data set used in this work,  $N = 22$ . For each parameter of the selected ECM, an initial guess was required. To ensure that the goodness-of-fit was as high as possible, for each cell, the EIS sample collected at Days = 0 and SoC = 0% was fit to a set of default initial guesses. For each subsequent EIS sample, the fitted parameter values obtained by the model for the initial EIS sample were used as initial guesses.

To effectively fit the EIS data to each ECM, all data in the positive real region (above the real axis) and the three highest points in the negative real region (below the real axis) were used.

### 3.2 Model Selection

The most parsimonious ECM was identified by calculating the root mean square error (RMSE) and information criteria values for each model. The RMSE was determined as

follows:

211

$$\text{RMSE}_{\text{ECM}} = \sqrt{\frac{1}{N} \sum_{n=1}^N ([Z'_{\text{EIS}}(\omega_n) - Z'_{\text{ECM}}(\omega_n; \Theta)]^2 + [Z''_{\text{EIS}}(\omega_n) - Z''_{\text{ECM}}(\omega_n; \Theta)]^2)} \quad (3)$$

where  $\Theta$  is the set of parameters for each ECM (with  $k$  as the number of parameters): 212

$$\left[ \begin{array}{l} \Theta_{1\text{W Randles}} = [L_0 R_0 R_1 C_1 R_2 C_2 W_0] \\ \Theta_{2\text{W Randles}} = [L_0 R_0 R_1 C_1 R_2 C_2 W_0 W_1] \\ \Theta_{1\text{ZARC}} = [L_0 R_0 R_1 \tau_1 \alpha_1 W_0] \\ \Theta_{2\text{ZARC}} = [L_0 R_0 R_1 \tau_1 \alpha_1 R_2 \tau_2 \alpha_2 W_0] \\ \Theta_{3\text{ZARC}} = [L_0 R_0 R_1 \tau_1 \alpha_1 R_2 \tau_2 \alpha_2 R_3 \tau_3 \alpha_3 W_0] \\ \Theta_{2\text{TL}} = [L_0 R_0 R_{\text{ion1}} Q_{s1} \alpha_1 R_{\text{ion2}} Q_{s2} \alpha_2 W_0] \end{array} \right] \quad \begin{array}{l} k = 7 \\ k = 8 \\ k = 6 \\ k = 9 \\ k = 12 \\ k = 9 \end{array}$$

213

Information criteria were also calculated to estimate the efficiency of each model by maximizing model fit and minimizing model complexity. Two types of information criteria were used: Akaike information criterion (AIC) and Bayesian information criterion (BIC). AIC uses a fixed penalty while BIC uses a penalty that increases logarithmically with sample size  $N$ . As a result, both information criteria favor the most parsimonious model; however, BIC favors simpler models more strongly than AIC. For both information criteria, a lower value indicates a more parsimonious model. The methods of calculation for each information criterion [71, 72] are defined as:

$$\text{AIC}_{\text{ECM}} = 2k + N \ln(\text{RMSE}_{\text{ECM}}^2) \quad (4)$$

214

$$\text{BIC}_{\text{ECM}} = k \ln(2k + N) + 2k + N \ln(\text{RMSE}_{\text{ECM}}^2). \quad (5)$$

215

### 3.3 SoH Estimation 223

While purely data-driven SoH estimation methods, such as GPR, support vector machines (SVMs), and CNNs, can achieve high accuracy, they often lack physical interpretability and may overfit limited datasets. On the other hand, integrating physics-based models with machine learning may constrain predictions to electrochemically explainable relationships. Our hybrid approach combines the predictive strength of machine learning with the interpretability of physics-informed features. Therefore, in our work, the estimation of SoH is carried out using the parameters of the most parsimonious model as input features for four supervised models: decision tree, random forest, gradient boosting, and linear regression. For each supervised regression model, three different model architectures were tested, as defined in Figure 4. Architecture 1 attempts to use all parameters of the most parsimonious model and SoC values as the input features. Architecture 2 attempts to use the single parameter that trends most with SoH (as determined by the SHAP analysis, explained in detail below) as the only input feature to determine if it is sufficient for

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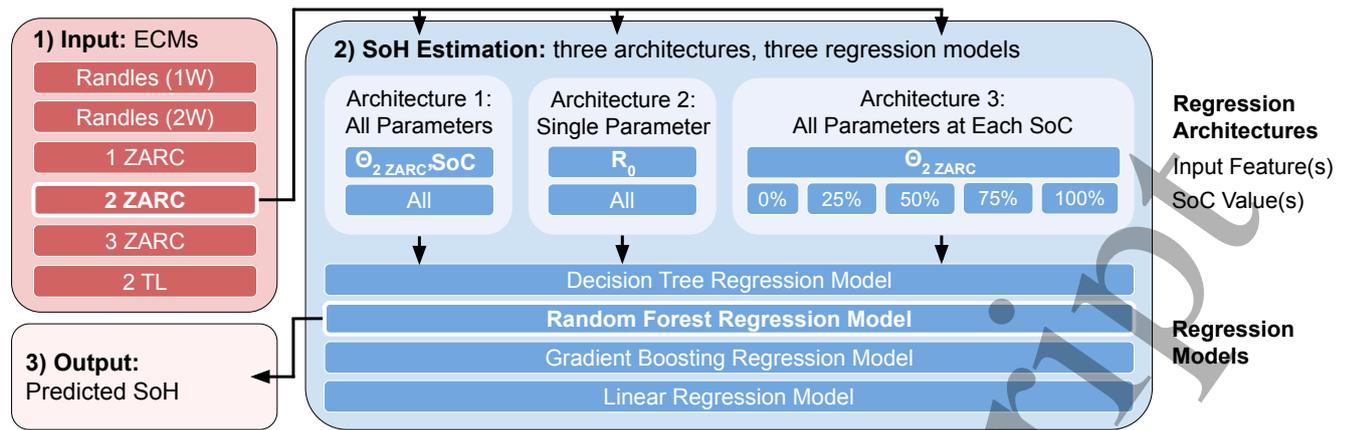


Figure 4: The three model architectures tested in the context of the entire model.

accurate SoH estimation. Finally, Architecture 3 attempts six different models, each using all parameters of the most parsimonious ECM as input features, at each SoC to determine if SoH estimation is more or less accurate at any single SoC.

The decision tree model was performed using `DecisionTreeRegressor` from the `sklearn.tree` module while the random forest and gradient boosting models used `RandomForestRegressor` and `GradientBoostingRegressor`, respectively, from the `sklearn.ensemble` module. The linear regression model was performed using `LinearRegression` from the `sklearn.linear_model` module. Each of the four regression models were evaluated using three different train/test splits: 70%/30%, 80%/20%, and 90%/10% of the entire dataset. The regression models were also evaluated at 21 random states, from 30 to 50, meaning that the models were trained and tested using initial random seeds within this range to assess performance consistency across varying random initializations. A random state controls the random subsampling of data for each estimator, or tree, as well as the random selection of features for splitting. In addition, the random forest and gradient boosting regression models were evaluated with 100 estimators.

For each regression model, the average RMSE was calculated across all train/test splits and random states. The RMSE of each regression model is defined as:

$$\text{RMSE}_{\text{Regression}} = \sqrt{\frac{1}{N} \sum_{n=1}^N (\text{SoH}_{\text{EIS}} - \text{SoH}_{2 \text{ ZARC}})^2}. \quad (6)$$

After the regression model with the lowest RMSE was selected, a follow-up SHapley Additive exPlanations (SHAP) multicollinearity correlation analysis was performed. SHAP is used to determine which features — including SoC and the parameters of the selected ECM — are most important in SoH estimation. The SHAP-explained impact of a feature is defined as:

$$\phi_i(f, x) = \sum_{z' \subseteq x'} \frac{|z'|!(M - |z'| - 1)!}{M!} [f_x(z') - f_x(z' \setminus i)], \quad (7)$$

where  $\phi_i(f, x)$  is the SHAP value for feature  $i$ ,  $f$  is the prediction model (e.g. decision tree, random forest, gradient boosting, or linear regression model),  $x$  is the input instance being explained (vector of feature values),  $x' \in \{0, 1\}^M$  is a binary vector indicating which features are included,  $z' \subseteq x'$  is a subset of included features,  $M$  is the total number of input features,  $|z'|$  is the number of features in the subset,  $f_x(z')$  is the model output when only the features in  $z'$  are known, and  $f_x(z' \setminus i)$  is the model output when feature  $i$  is removed from the subset [73]. A higher mean absolute SHAP value indicates a stronger correlation with SoH.

As an example of the SHAP analysis, suppose our SoH regressor uses just two features:  $R_0$  and SoC. For one EIS sample, the model's outputs for each subset of features are: with no features SoH = 0.82, with  $R_0$  only SoH = 0.78, with SoC only SoH = 0.88, and with both SoH = 0.84. Since  $M = 2$ , the SHAP weight on each marginal contribution is  $\frac{1}{2}$ . Thus,

$$\phi_{R_0} = \frac{1}{2}(0.78 - 0.82) + \frac{1}{2}(0.84 - 0.88) \quad (8)$$

and

$$\phi_{SoC} = \frac{1}{2}(0.88 - 0.82) + \frac{1}{2}(0.84 - 0.78). \quad (9)$$

For each of the models used in this section, SHAP values are calculated for each of the respective input features with respect to SoH to determine the most impactful of these features. Architecture 2 in the following section uses only the most impactful feature, that is, the feature with the highest SHAP value, as the sole input.

## 4. Results

### 4.1 Assessment of ECMs

The EIS data from all 22 cells at different ages and SoCs were fit to the six ECMs. The RMSEs and individual parameter values from each fit were then extracted and averaged for each ECM. Table 2 shows the average value and standard deviation for each parameter. The best and worst fits for each ECM are shown in Figure 5. This work finds that the 2 ZARC model was the most accurate in fitting the data, as demonstrated by both the RMSE and information criterion evaluation results.

Parameter	Average Value	Average Standard Deviation
$L_0$	$5.7988 \times 10^{-8} [H]$	$2.03 \times 10^{-8}$
$R_0$	$1.3525 \times 10^{-3} [\Omega]$	$2.37 \times 10^{-4}$
$R_1$	$3.9600 \times 10^{-4} [\Omega]$	$5.75 \times 10^{-4}$
$\tau_1$	$1.3900 \times 10^{-2} [s]$	$2.05 \times 10^{-1}$
$\alpha_1$	$7.9419 \times 10^{-1}$	$2.14 \times 10^{-1}$
$R_2$	$6.8001 \times 10^{-1} [\Omega]$	$1.29 \times 10^1$
$\tau_2$	$4.1100 \times 10^4 [s]$	$6.92 \times 10^5$
$\alpha_2$	$8.7071 \times 10^{-1}$	$1.27 \times 10^{-1}$
$W_0$	$1.6603 \times 10^{-4} [\Omega/s^{1/2}]$	$1.95 \times 10^{-4}$

Table 2: Average values

of extracted parameter and standard deviation values from 2 ZARC ECM fits of all EIS samples.

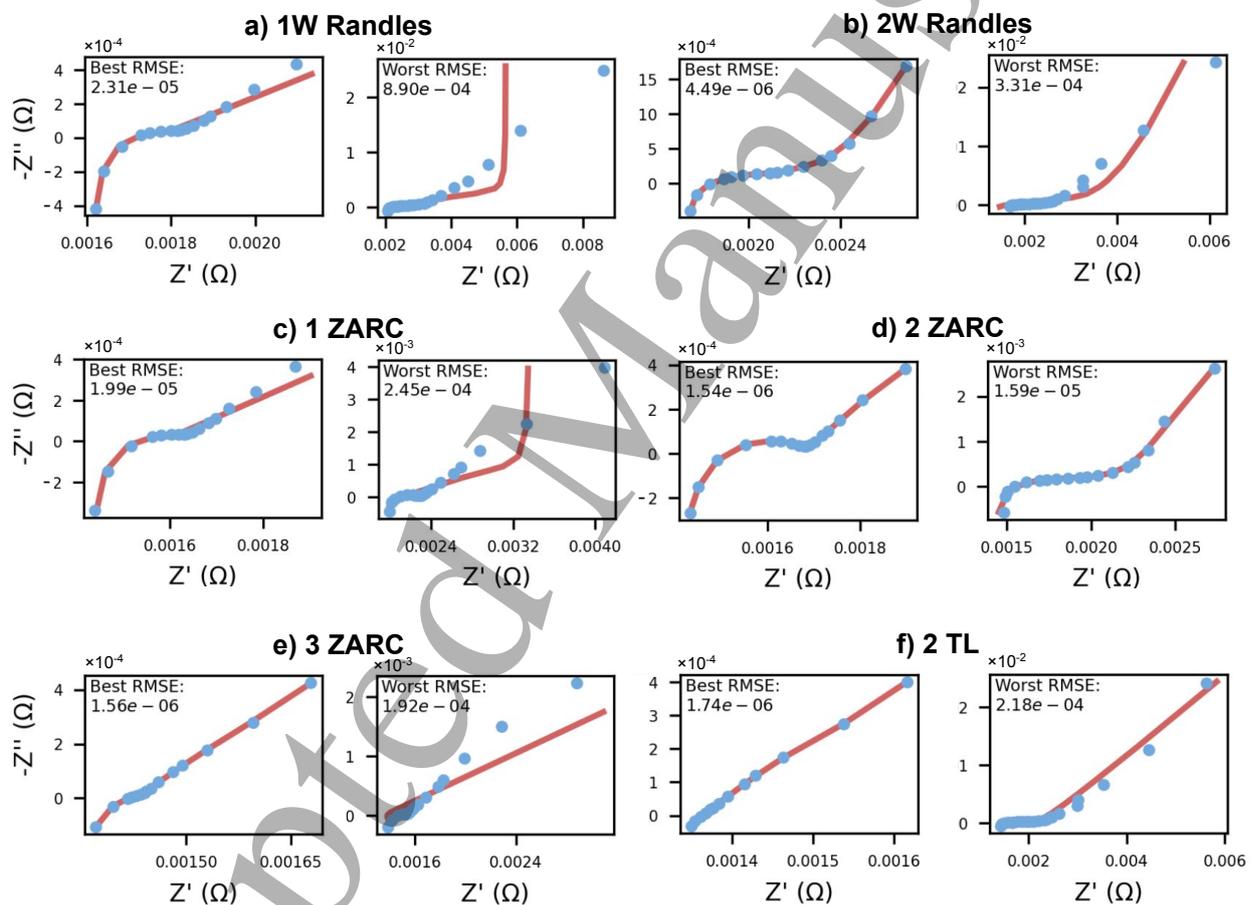


Figure 5: The best fit, corresponding to the lowest RMSE, and worst fit, corresponding to the highest RMSE, across all cells for each ECM: a) 1W Randles, b) 2W Randles, c) 1 ZARC, d) 2 ZARC, e) 3 ZARC, and f) 2 TL.

First, the average RMSE of the 2 ZARC circuit was found to be significantly lower than that of the 1W Randles, 2W Randles, 1 ZARC, and 2 TL circuits, and slightly lower than that of the 3 ZARC circuit, as shown in Figure 6a. The median RMSE of each ECM also presents similar trends, demonstrating that the performance of the 2 ZARC

1  
2 circuit is consistent across most cells and not driven by outliers. This indicates that 289  
3 the ZARC element successfully captured the non-ideal behavior of the capacitor in both 290  
4 RC components. This improvement stems from the inclusion of a CPE in the ZARC, 291  
5 which generalizes ideal capacitive behavior by introducing a fractional-order impedance 292  
6 ( $\alpha$ ); unlike a standard RC element, the CPE can model the distributed and non-ideal 293  
7 electrochemical processes that arise from surface roughness, porosity, and inhomogeneities 294  
8 within the electrode [74]. In addition, because the ZARC component aimed to specifically 295  
9 model the behavior of the double-layer capacitance within a cell, which is represented by 296  
10 the second ZARC element, the 2 ZARC circuit further performed better than the 1 ZARC 297  
11 circuit. 298

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17 Second, the 2 ZARC circuit was found to have the most optimal balance of accuracy 299  
18 and efficiency. Both the AIC and BIC tests found that the 2 ZARC had the lowest value 300  
19 among all six ECMs, indicating that it was the most parsimonious, as shown in Figure 301  
20 6a. Similar to the RMSE results, the 1W Randles, 2W Randles, 1 ZARC, and 2 TL 302  
21 circuits had significantly higher AIC and BIC values than both the 2 ZARC and 3 ZARC 303  
22 circuits. For real-world conditions, this means that the 2 ZARC ECM would likely require 304  
23 comparatively less computational time and energy to simulate data in general. 305

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28 Across the five different SoCs, the average RMSE of each ECM's fit was relatively 306  
29 equal, except for SoC = 0%, which had a significantly greater average RMSE. This can 307  
30 be attributed to the Warburg tail component at SOC = 0%, which is much harder to fit 308  
31 because it is not precisely 45 degrees from the x-axis. An example of this is the best fit in 309  
32 Figure 5. 310  
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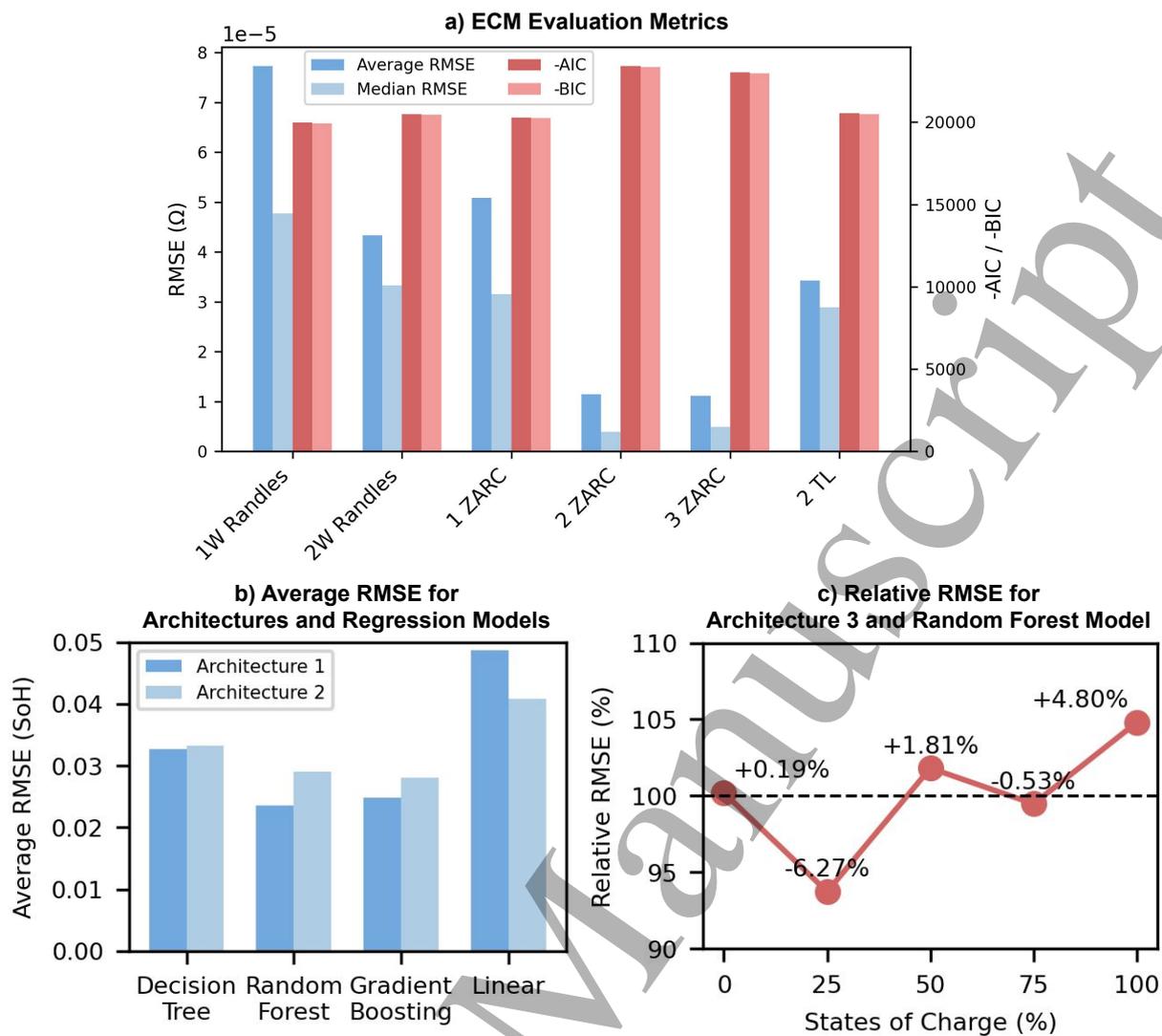


Figure 6: Across all cells, (a) average RMSE, median RMSE, -AIC, and -BIC comparison across ECMs, (where lower AIC and BIC values, shown as taller columns, indicate better model fit penalized by complexity), (b) the average RMSEs of Architectures 1 and 2 with decision tree, random forest, gradient boosting, and linear regression models, and (c) relative RMSEs of Architecture 3 with random forest regression model, where 100% relative RMSE indicates the average RMSE across all SoCs.

## 4.2 Evaluation of Regression Models

Next, the decision tree, random forest, and gradient boosting, and linear regression models were each executed under three distinct model architectures.

The regression models were first performed under Architecture 1, which used all of the parameters extracted from the 2 ZARC model as well as SoH, as input features. The results from Architecture 1, as shown in Figure 6b, suggested that the random forest estimated SoH best. This specific architecture-model pipeline achieved an average RMSE of 0.0236 (where SoH is normalized to range from 0 to 1), a 96.4% accuracy within a 95% confidence interval, and a 0.8762  $R^2$  value (indicating that 87.62% of the variance in the actual SoH values is explained by the model's estimation) for the best train/test/random

1 state split of 90% train/0.05% test data and random state 42 (Figure 7a). This is consistent 321  
2 with the fact that the random forest model is generally more robust to noise and less prone 322  
3 to overfitting, as it averages estimations across several trees. The density plot (Figure 7b) 323  
4 of RMSE across different train/test splits indicates that the greater training proportions 324  
5 tend to result in relatively lower RMSE distributions than smaller training proportions. 325

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10 Work utilizing other approaches on the same data has been performed [46]. While 326  
11 our work achieves an average RMSE of 2.36% for the hybrid ECM-Random Forest model 327  
12 and 4.86% for the hybrid ECM-Linear Regression model, an alternative approach using 328  
13 distribution of relaxation times (DRT) achieves an average RMSE of 1.69% for the hybrid 329  
14 DRT-Long Short-Term Memory model and 4.89% for the hybrid DRT-Linear Regression 330  
15 model. 331

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18 While the random forest model is typically less interpretable than a single decision 332  
19 tree, the SHAP feature importance analysis provided a critical method for understanding 333  
20 the impact that individual features may have on the accuracy of SoH estimation. The 334  
21 SHAP feature importance analysis found that the ohmic resistance feature ( $R_0$ ) is most 335  
22 significant in SoH estimation (Figure 7c) and indeed had a negative correlation with SoH 336  
23 (Figure 7d). A higher  $R_0$  may indicate greater degradation, such as electrode cracking. 337  
24 These aging effects reduce ionic and electronic conductivity within the cell, resulting in a 338  
25 lower available capacity [75]. Following  $R_0$ , the  $\tau_1$  component showed a weaker, yet still 339  
26 valuable, feature importance with SoH. This is consistent with the fact that slower speeds 340  
27 of electrochemical reactions may represent slower charge transfer kinetics and increased 341  
28 diffusion resistance. A thicker insulating layer induces rapid performance degradation and 342  
29 longer ion diffusion times, resulting in a lower SoH [76]. 343

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36 To evaluate whether  $R_0$  alone could predict SoH, Architecture 2 tested models using 344  
37 only this feature. The random forest model performed worse, with an average RMSE of 345  
38 0.0290 (Figure 6b). Similar results were observed across the other three models, where 346  
39 Architecture 2 generally underperformed relative to Architecture 1. The only exception was 347  
40 linear regression, which improved slightly, likely because excluding less relevant features 348  
41 reduced sensitivity to multicollinearity. Overall, these results indicate that  $R_0$  by itself is 349  
42 insufficient for accurate SoH estimation, and additional features are necessary. 350

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47 Finally, Architecture 3 was performed to evaluate the accuracy of the random forest 351  
48 model on SoC-specific data. The data was separated by SoC into five testing sets. All 352  
49 parameters of the 2 ZARC model were used as features. The results of Architecture 3 353  
50 suggest that the random forest model performs relatively similarly across all SoCs (Figure 354  
51 6c). While the model at SoC = 25% did exhibit the lowest RMSE, its RMSE was only 355  
52 6.27% less than that of the average RMSE. The model at SoC = 100% exhibited the 356  
53 highest RMSE, which was 4.80% greater than the average RMSE. 357

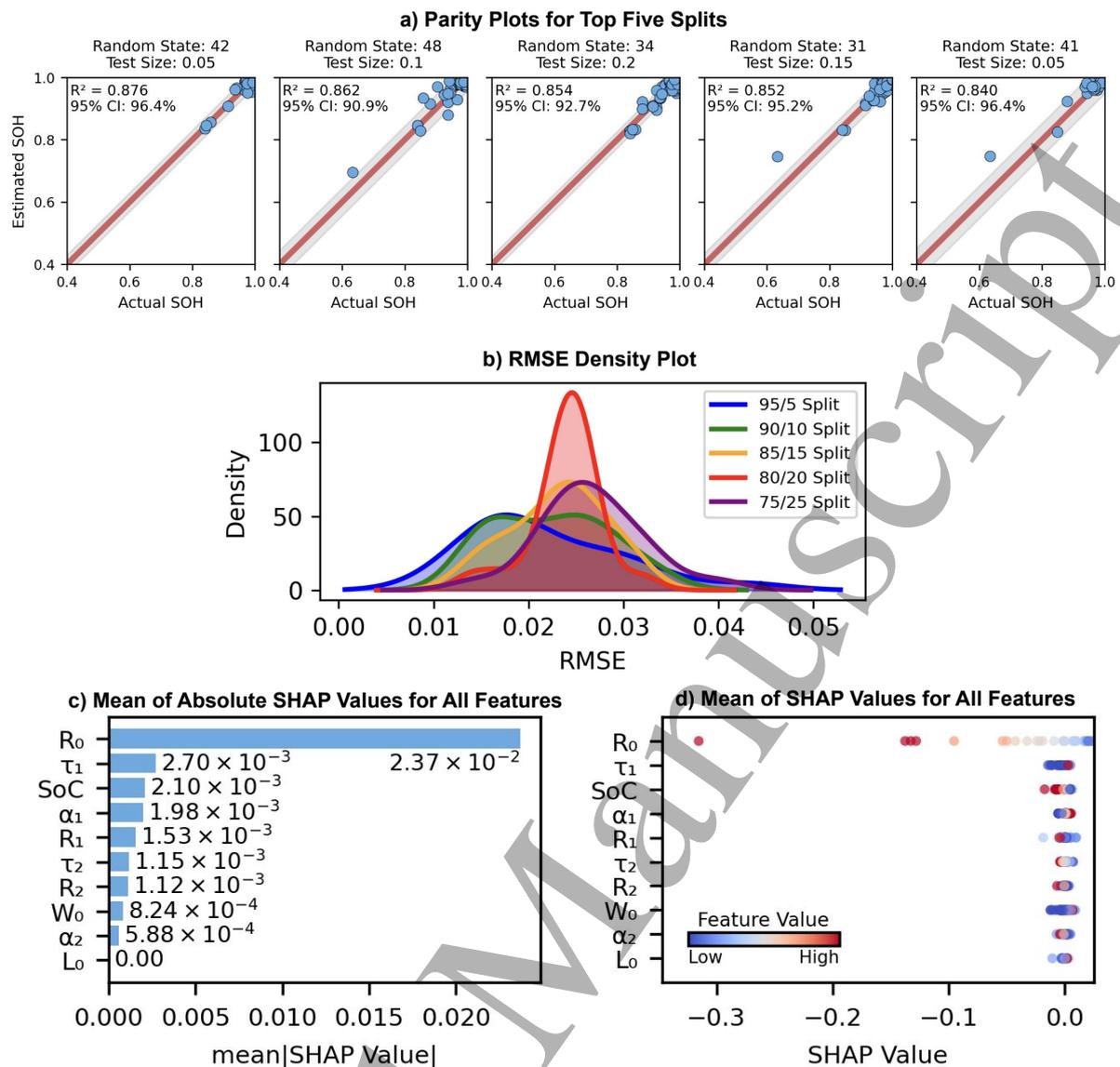


Figure 7: For the random forest regression model under Architecture 1: a) parity plots for the top five splits (highest  $R^2$  across all combinations of random states 30–50 and select test sizes 0.05, 0.10, 0.15, 0.20, and 0.25), where "95% CI" indicates the percentage of predictions falling within 95% confidence interval (gray area), b) density plot of RMSE for the selected test sizes, each evaluated 100 times, c) absolute mean SHAP values for all features, and d) mean SHAP values colored by feature value.

## 5. Conclusions

This study provides the groundwork for a comprehensive and explainable estimation tool that fits EIS data to multiple ECMs, uses comparative analysis to determine the most accurate model, and ultimately uses the parameters from this model to estimate SoH. The results notably show that the ohmic resistance of a cell's electrolyte, electrodes, and current collectors is a critical factor in estimating its health. However, the inclusion of additional parameters is also vital for accurate estimation.

Furthermore, two regression model architectures, one that includes all parameters of

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2 the most accurate and efficient ECM and one that includes the ohmic resistance element, 366  
3 prove that the latter is not sufficient for accurate battery health estimation. A third model 367  
4 architecture finds that SoH estimation is relatively consistent across all states of charge of 368  
5 the battery. 369  
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8 Although our work is validated on 5 Ah NMC/graphite cells, the proposed framework, 370  
9 which includes SHAP-based feature importance analysis, can be applied to other Li-ion 371  
10 chemistries. Because cells of different chemistries are all characterized by degradation 372  
11 mechanisms such as SEI growth, loss of active material, and electrolyte decomposition, 373  
12 the application of the model framework will likely remain similar, even though the specific 374  
13 values and types of extracted ECM parameters may differ due to material properties. 375  
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16 Finally, limitations include that EIS is not widely feasible for continuous monitoring in 376  
17 commercial BMSs due to hardware limitations and measurement constraints. Future work 377  
18 could focus on current-voltage pulse responses or drive-cycle data, which are more readily 378  
19 available onboard. The deployment of on-board EIS technologies for diagnostics would 379  
20 also benefit from the pipeline presented in this work [77]. 380  
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## 7. Appendix

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It is important to note that a fourth model architecture was tested, which had the same input features as Architecture 1, except for  $R_0$ . This model architecture performed worse than both Architectures 1 and 2, with an average RMSE of 0.0371 for the Random Forest model. Table 4 is a look-up table for the fitted parameters of the 2 ZARC model and their associated SoH.

Cycle Day	SoC	$L_0$	$R_0$	$R_1$	$\tau_1$	$\alpha_1$	$R_2$	$\tau_1$	$\alpha_1$	$W_0$	SoH
0	0	6.11E-08	1.32E-03	3.00E-04	9.17E-04	8.40E-01	8.82E-01	2.04E+05	8.92E-01	1.98E-04	1.000
10	0	5.51E-08	1.29E-03	3.94E-04	3.38E-02	7.80E-01	4.66E-02	9.03E+01	8.79E-01	1.55E-04	0.986
20	0	5.48E-08	1.32E-03	3.98E-04	2.87E-03	7.90E-01	1.09E-01	3.85E+04	8.51E-01	1.43E-04	0.978
40	0	5.23E-08	1.35E-03	4.25E-04	3.65E-02	7.91E-01	3.41E-02	2.55E+03	8.57E-01	1.55E-04	0.961
90	0	5.80E-08	1.46E-03	4.13E-04	7.81E-03	7.88E-01	1.63E-02	6.90E+01	8.69E-01	1.70E-04	0.918
0	25	6.11E-08	1.32E-03	3.00E-04	9.17E-04	8.40E-01	8.82E-01	2.04E+05	8.92E-01	1.98E-04	1.000
10	25	5.51E-08	1.29E-03	3.94E-04	3.38E-02	7.80E-01	4.66E-02	9.03E+01	8.79E-01	1.55E-04	0.986
20	25	5.48E-08	1.32E-03	3.98E-04	2.87E-03	7.90E-01	1.09E-01	3.85E+04	8.51E-01	1.43E-04	0.978
40	25	5.23E-08	1.35E-03	4.25E-04	3.65E-02	7.91E-01	3.41E-02	2.55E+03	8.57E-01	1.55E-04	0.961
90	25	5.80E-08	1.46E-03	4.13E-04	7.81E-03	7.88E-01	1.63E-02	6.90E+01	8.69E-01	1.70E-04	0.918
0	50	6.11E-08	1.32E-03	3.00E-04	9.17E-04	8.40E-01	8.82E-01	2.04E+05	8.92E-01	1.98E-04	1.000
10	50	5.51E-08	1.29E-03	3.94E-04	3.38E-02	7.80E-01	4.66E-02	9.03E+01	8.79E-01	1.55E-04	0.986
20	50	5.48E-08	1.32E-03	3.98E-04	2.87E-03	7.90E-01	1.09E-01	3.85E+04	8.51E-01	1.43E-04	0.978
40	50	5.23E-08	1.35E-03	4.25E-04	3.65E-02	7.91E-01	3.41E-02	2.55E+03	8.57E-01	1.55E-04	0.961
90	50	5.80E-08	1.46E-03	4.13E-04	7.81E-03	7.88E-01	1.63E-02	6.90E+01	8.69E-01	1.70E-04	0.918
0	75	6.11E-08	1.32E-03	3.00E-04	9.17E-04	8.40E-01	8.82E-01	2.04E+05	8.92E-01	1.98E-04	1.000
10	75	5.51E-08	1.29E-03	3.94E-04	3.38E-02	7.80E-01	4.66E-02	9.03E+01	8.79E-01	1.55E-04	0.986
20	75	5.48E-08	1.32E-03	3.98E-04	2.87E-03	7.90E-01	1.09E-01	3.85E+04	8.51E-01	1.43E-04	0.978
40	75	5.23E-08	1.35E-03	4.25E-04	3.65E-02	7.91E-01	3.41E-02	2.55E+03	8.57E-01	1.55E-04	0.961
90	75	5.80E-08	1.46E-03	4.13E-04	7.81E-03	7.88E-01	1.63E-02	6.90E+01	8.69E-01	1.70E-04	0.918
0	100	6.11E-08	1.32E-03	3.00E-04	9.17E-04	8.40E-01	8.82E-01	2.04E+05	8.92E-01	1.98E-04	1.000
10	100	5.51E-08	1.29E-03	3.94E-04	3.38E-02	7.80E-01	4.66E-02	9.03E+01	8.79E-01	1.55E-04	0.986
20	100	5.48E-08	1.32E-03	3.98E-04	2.87E-03	7.90E-01	1.09E-01	3.85E+04	8.51E-01	1.43E-04	0.978
40	100	5.23E-08	1.35E-03	4.25E-04	3.65E-02	7.91E-01	3.41E-02	2.55E+03	8.57E-01	1.55E-04	0.961
90	100	5.80E-08	1.46E-03	4.13E-04	7.81E-03	7.88E-01	1.63E-02	6.90E+01	8.69E-01	1.70E-04	0.918

Table 3: Look-up table for the fitted parameters of the 2 ZARC model and the associated SoH at five different SoCs and five different days of aging.

In addition, to better understand how feature importance impacts estimation accuracy at certain SoHs, Figure 8 visualizes the actual versus estimated SoH across top-performing splits.

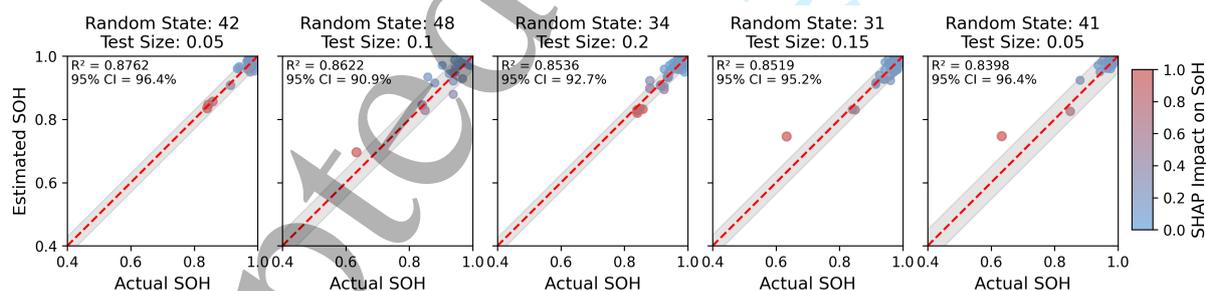


Figure 8: Parity plots for the top five splits (highest  $R^2$  across all combinations of random states 30–50 and select test sizes 0.05, 0.10, 0.15, 0.20, and 0.25), colored by normalized SHAP impact on SoH to represent the total contribution of all features to the estimated SoH.